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# Fluctuations of the Mutual Information in Large Distributed Antenna Systems with Colored Noise

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**Abstract**—This paper studies the fluctuations of the mutual information of a class of multiple-input multiple-output (MIMO) channels with arbitrary correlated noise in the large system limit. We provide a deterministic approximation of the ergodic mutual information, which is asymptotically accurate as the number of antennas grows, and study its fluctuations around this value under the form of a central limit theorem (CLT). This result can be used to predict the outage capacity for slow fading channels. The channel model considered in this contribution has a particular application in the context of distributed antenna or network MIMO systems where the path loss between any pair of transmit and receive antennas has a different value. As shown by simulations, the approximations are very accurate for channels of small dimensions.

## I. INTRODUCTION

Consider a wireless communication channel between  $n$  single antenna transmitters and a receiver equipped with  $N$  antennas. Let  $\mathbf{H} \in \mathbb{C}^{N \times n}$  be the channel matrix representing the complex channel gains from the transmitters to the receiver. The receive vector  $\mathbf{y} \in \mathbb{C}^N$  at a given time instant reads

$$\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{z} \quad (1)$$

where  $\mathbf{x} \in \mathbb{C}^n$  is the vector of the transmitted signals and  $\mathbf{z} \in \mathbb{C}^N$  is a vector of complex Gaussian noise with covariance matrix  $\mathbb{E}[\mathbf{z}\mathbf{z}^H] = \rho\mathbf{I}_N + \mathbf{A}\mathbf{A}^H$ . Typically,  $\rho\mathbf{I}_N$  represents an uncorrelated thermal noise component with power  $\rho$  while  $\mathbf{A}\mathbf{A}^H$  accounts for a source of correlated interference whose covariance matrix has the non-negative square root  $\mathbf{A} \in \mathbb{C}^{N \times m}$ . The  $(i, j)$ -entry  $h_{ij}$  of the channel matrix  $\mathbf{H}$  is modeled as

$$h_{ij} = \frac{\sigma_{ij}}{\sqrt{n}} w_{ij}$$

where  $(\sigma_{ij}^2)$  is a sequence of non-negative real numbers called a variance profile and the  $w_{ij}$  are independent, standard complex Gaussian random variables with zero mean and unit variance. For a complex Gaussian channel input vector  $\mathbf{x}$  with covariance matrix  $\mathbb{E}[\mathbf{x}\mathbf{x}^H] = \mathbf{I}_n$  and full channel knowledge at the receiver, the normalized ergodic capacity of the channel is given by  $I(\rho) = \mathbb{E}[Z(\rho)]$ , where

$$I(\rho) = \frac{1}{N} \log \det \left( \mathbf{I}_N + (\rho\mathbf{I} + \mathbf{A}\mathbf{A}^H)^{-1} \mathbf{H}\mathbf{H}^H \right) . \quad (2)$$

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The aim of this paper is to derive a deterministic approximation  $V(\rho)$  of  $I(\rho)$  and to study the fluctuations of the random variable  $N(\mathcal{I}(\rho) - V(\rho))$  in the large system limit, i.e., for  $N, n \rightarrow \infty$  at the same pace. More precisely, the notation  $N, n \rightarrow \infty$  will refer in the sequel to the two following conditions on  $n$ ,  $N$  and the number of columns  $m$  of the matrix  $\mathbf{A}$ :

$$\begin{aligned} 0 &< \liminf_{n \rightarrow \infty} \frac{N}{n} \leq \limsup_{n \rightarrow \infty} \frac{N}{n} < \infty \\ 0 &\leq \limsup_{n \rightarrow \infty} \frac{m}{n} < \infty . \end{aligned} \quad (3)$$

A well known result of random matrix theory states that the empirical eigenvalue distribution of the Gram matrix  $\mathbf{H}\mathbf{H}^H$  converges weakly to a limit distribution function when the elements of  $\mathbf{H}$  are independent and identically distributed (i.i.d), i.e.,  $\sigma_{ij}^2 = 1$  [1]. This fact leads also to the convergence of  $\mathcal{I}(\rho)$  to a deterministic limit which can be given in closed form [2]. Similar results could be established for more complicated models such as the Kronecker model  $\mathbf{H} = \mathbf{\Phi}_R \mathbf{W} \mathbf{\Phi}_T$  [3], [4], where  $\mathbf{W}$  is a  $N \times n$  standard Gaussian matrix and  $\mathbf{\Phi}_R$  and  $\mathbf{\Phi}_T$  are  $N \times N$  and  $n \times n$  matrices capturing the effects of transmit and receive antenna correlation, sums of matrices  $\mathbf{H}_k$  each having a Kronecker variance structure [5] and also non-centered channel matrices with a variance profile [6]. These works provide deterministic approximations  $V(\rho)$  of  $I(\rho)$ , only depending on  $N, n$  and the distribution of  $\mathbf{H}$ , in the sense that  $I(\rho) - V(\rho) \rightarrow 0$  for  $n, N \rightarrow \infty$  while satisfying (3). Apart from some special cases, the function  $V(\rho)$  is rarely available in closed-form and requires to solve a set of implicit equations. However, their computation is in general much less complex than the capacity evaluation by Monte Carlo simulations. Moreover, the deterministic approximations have been shown by simulations to yield very accurate results for small channel dimensions with as little as two transmit and receive antennas. Recently, the fluctuations of the mutual information have been studied under the form of central limit theorems (CLTs). One is generally interested in obtaining results of the form:

$$\frac{N^\alpha}{\Theta} (\mathcal{I}(\rho) - V(\rho)) \rightarrow \mathcal{N}(0, 1) \quad (4)$$

in distribution, where  $\alpha$  is a measure of the convergence speed and  $\Theta^2$  determines the variance. In a slow fading

scenario, these results allow to approximate the outage probability, i.e.,  $\Pr(N\mathcal{I}(\rho) \leq R)$ , for a given desired target rate  $R$ . A CLT for channel matrices with left-sided correlation was established in [7] and the more general case of a variance profile addressed in [8]. Also the fluctuations of the mutual information in the presence of correlated interference and noise under the Kronecker model were studied in [4] relying on the replica method.

The novelty of the results derived in this paper in contrast to [8] is the consideration of arbitrary correlated Gaussian noise whose covariance matrix can be written in the form  $\rho\mathbf{I}_N + \mathbf{A}\mathbf{A}^H$ . This model is more general than the particular case where the interference term can be written in the form  $\mathbf{H}_I\mathbf{x}_I$ , where  $\mathbf{H}_I$  is a random matrix which follows the same statistical model as the channel matrix  $\mathbf{H}$  and  $\mathbf{x}_I$  is a standard complex Gaussian vector. Here, the mutual information can be decomposed into two terms without interference, i.e.,  $\mathcal{I}(\rho) = \frac{1}{N} \log \det(\rho\mathbf{I}_N + \mathbf{H}_I\mathbf{H}_I^H + \mathbf{H}\mathbf{H}^H) - \frac{1}{N} \log \det(\rho\mathbf{I}_N + \mathbf{H}_I\mathbf{H}_I^H)$ , where the first can be seen as the mutual information of the compound channel  $[\mathbf{H}\mathbf{H}_I]$  and the second as the mutual information of the interference channel  $\mathbf{H}_I$ . Note that both matrices  $\mathbf{H}$  and  $\mathbf{H}_I$  are considered random while  $\mathbf{A}$  in our model is assumed to be deterministic.

The channel model considered in this work has a particular application in the context of distributed antenna or network MIMO systems where the signals received at several spatially separated antennas are jointly processed to provide macro diversity. For more details on this topic we refer the reader to the comprehensive surveys [9], [10]. More precisely, the channel in (1) can be seen as a multiple access channel (MAC) with macro diversity where the value of a particular  $\sigma_{ij}^2$  represents the inverse path loss between the  $j$ th transmitter and the  $i$ th receiving antenna. For example, assuming a log-distance path loss model, we have  $\sigma_{ij}^2 = d_{ij}^{-\beta}$ , where  $d_{ij}$  is the distance between transmitter  $j$  and receive antenna  $i$  and  $\beta$  is the path loss exponent whose value lies usually in the range from 2 to 5 depending on the radio environment. The application of random matrix theory to the study of multi-cellular networks is not new. The sum-capacity scaling of large cooperative cellular networks has been studied in [11] and the downlink of large multi-cell systems with optimal power allocation and user scheduling was considered in [12]. However, both works build upon the assumption that the inter-cell interference is of the form  $\mathbf{H}_I\mathbf{x}_I$  as discussed above.

The remainder of the paper is structured as follows. A first order result in form of a deterministic equivalent of the normalized mutual information is derived in Section II. That is, we find a deterministic function  $V(\rho)$  such that  $I(\rho) - V(\rho) \rightarrow 0$ , for  $N, n \rightarrow \infty$ . The fluctuations of the random variable  $N(\mathcal{I}(\rho) - V(\rho))$  are studied in Section III where we establish a CLT of the form (4) and provide an explicit expression for the asymptotic variance  $\Theta^2$ . Numerical results are presented in Section IV which corroborate the theoretical results and demonstrate their applicability to channel matrices of even small dimensions. Section V concludes the paper.

## II. DETERMINISTIC APPROXIMATION OF $I(\rho)$

Recall that  $\mathbf{H}$  is a  $N \times n$  matrix,  $\mathbf{A}$  is  $N \times m$ . The aim of this section is to propose a deterministic equivalent to the normalized ergodic mutual information

$$I(\rho) = \frac{1}{N} \mathbb{E} [\log \det(\rho\mathbf{I}_N + \mathbf{A}\mathbf{A}^H + \mathbf{H}\mathbf{H}^H)] - \frac{1}{N} \log \det(\rho\mathbf{I}_N + \mathbf{A}\mathbf{A}^H)$$

as  $N, n \rightarrow \infty$ . Consider the following technical assumptions and assume that  $n, N$  and  $m$  satisfy (3):

**A 1:** Consider a family of non-negative real numbers  $(\sigma_{ij}^{(n)}; 1 \leq i \leq N; 1 \leq j \leq n)$ , then there exists a non-negative real number  $\sigma_{\max}$  such that:

$$\sup_{i \leq N, j \leq n, n \geq 1} \sigma_{ij}^{(n)} \leq \sigma_{\max} < \infty.$$

**A 2:** Consider a family  $(\mathbf{A}_{N,m}; N \geq 1, m \geq 1)$  of  $N \times m$  matrices and denote by  $\|\mathbf{A}_{n,m}\|$  the spectral norm of matrix  $\mathbf{A}_{n,m}$ , then there exists a non-negative real number  $a_{\max}$  such that:

$$\sup_{N,m} \|\mathbf{A}_{N,m}\| \leq a_{\max} < \infty.$$

**A 3:** Consider a family of non-negative real numbers  $(\sigma_{ij}^{(n)}; 1 \leq i \leq N; 1 \leq j \leq n)$ . Then there exists a positive number  $\sigma_{\min}$  such that:

$$\liminf_{n \geq 1} \min_{1 \leq j \leq n} \frac{1}{n} \sum_{i=1}^N \sigma_{ij}^{(n)} \geq \sigma_{\min}^2.$$

In the sequel, we will drop the dependencies in  $N, n$  and  $m$  and simply write  $\sigma_{ij}$  and  $\mathbf{A}$  instead of  $\sigma_{ij}^{(n)}$  and  $\mathbf{A}_{N,m}$ . Consider the following diagonal  $N \times N$  matrices:

$$\mathbf{D}_j = \text{diag}(\sigma_{1j}^2, \dots, \sigma_{Nj}^2), \quad 1 \leq j \leq n. \quad (5)$$

Denote by  $\mathbb{C}_+ = \{z \in \mathbb{C} : \text{Im}(z) > 0\}$ , and by  $\mathcal{S}$  the class of functions  $f$  analytic over  $\mathbb{C}_+$ , such that  $f : \mathbb{C}_+ \rightarrow \mathbb{C}_+$  and  $\lim_{y \rightarrow 0} -iyf(iy) = 1$ , where  $\mathbf{i} = \sqrt{-1}$ .<sup>1</sup> We are now in position to state the first result of the paper:

*Theorem 1 (Deterministic Equivalent):* Assume that assumptions (A1) and (A2) hold true, then:

(i) The following equation:

$$\mathbf{T}(z) = \left( \mathbf{A}\mathbf{A}^H - z\mathbf{I}_N + \frac{1}{n} \sum_{j=1}^n \frac{\mathbf{D}_j}{1 + \frac{1}{n} \text{tr} \mathbf{D}_j \mathbf{T}(z)} \right)^{-1} \quad (6)$$

admits a unique solution  $\mathbf{T}(z)$  in the space of  $N \times N$  matrices such that there exists a  $N \times N$  matrix-valued measure<sup>2</sup>  $\mu$  such that:

$$\mathbf{T}(z) = \int_{\mathbb{R}^+} \frac{\mu(d\lambda)}{\lambda - z} \quad \text{where} \quad \mu(\mathbb{R}^+) = \mathbf{I}_N. \quad (7)$$

In particular,  $\frac{1}{N} \text{tr} \mathbf{T}(z) \in \mathcal{S}$ .

<sup>1</sup>Such functions are known to be Stieltjes transforms of probability measures over  $\mathbb{R}$  - see for instance [6, Proposition 2.2].

<sup>2</sup>For details, see for instance [6, Theorem 2.4].

(ii) Let  $\rho > 0$ . Denote  $\mathbf{T}_\rho = \mathbf{T}(-\rho)$  and consider the quantity:

$$\begin{aligned} V(\rho) = & -\frac{1}{N} \log \det (\mathbf{T}_\rho (\rho \mathbf{I}_N + \mathbf{A} \mathbf{A}^H)) \\ & + \frac{1}{N} \sum_{j=1}^n \log \left( 1 + \frac{1}{n} \text{tr} \mathbf{D}_j \mathbf{T}_\rho \right) \\ & - \frac{1}{Nn} \sum_{\substack{i=1, \dots, N \\ j=1, \dots, n}} \frac{\sigma_{ij}^2 \mathbf{T}_{ii}(-\rho)}{1 + \frac{1}{n} \text{tr} \mathbf{D}_j \mathbf{T}_\rho} . \end{aligned}$$

Then, the following holds true:

$$I(\rho) - V(\rho) \xrightarrow[N, n \rightarrow \infty]{} 0 .$$

*Proof:* Theorem 1 is essentially a consequence of Theorems 2.4 and 4.1 in [6]. The main idea is to cast the model  $\mathbf{H} \mathbf{H}^H + \mathbf{A} \mathbf{A}^H$  into an extended model which fits into the framework of [6]. Consider the  $N \times (n+m)$  matrices  $\mathbf{Z} = [\mathbf{H} \ \mathbf{0}_{N \times m}]$  and  $\mathbf{\Gamma} = [\mathbf{0}_{N \times n} \ \mathbf{A}]$ ; then  $(\mathbf{Z} + \mathbf{\Gamma})(\mathbf{Z} + \mathbf{\Gamma})^H = \mathbf{H} \mathbf{H}^H + \mathbf{A} \mathbf{A}^H$ , which is precisely the model under investigation. Introduce the following notations, for  $1 \leq i \leq N$  and  $1 \leq j \leq n+m$ :

$$\begin{aligned} \rho_{ij} &= \begin{cases} \sqrt{\frac{n+m}{n}} \times \sigma_{ij} & \text{if } j \leq n \\ 0 & \text{if } j \geq n+1 \end{cases} \\ \tilde{\Delta}_i &= \text{diag}(\rho_{ij}^2; 1 \leq j \leq n+m) \\ \Delta_j &= \text{diag}(\rho_{ij}^2; 1 \leq i \leq N) . \end{aligned}$$

Note that  $\Delta_j = \mathbf{0}_{N \times N}$  if  $j \geq n+1$ . We can now write down the equations associated to the model  $(\mathbf{Z} + \mathbf{\Gamma})(\mathbf{Z} + \mathbf{\Gamma})^H$  as given in [6, Theorem 2.4]. Let  $\Psi(z) = \text{diag}(\Psi_i(z), 1 \leq i \leq N)$  and  $\tilde{\Psi}(z) = \text{diag}(\tilde{\Psi}_j(z), 1 \leq j \leq n+m)$  where:

$$\begin{aligned} \Psi_i(z) &= -\frac{1}{z \left( 1 + \frac{1}{n+m} \text{tr} \tilde{\Delta}_i \tilde{\mathbf{T}}(z) \right)} \\ \tilde{\Psi}_j(z) &= -\frac{1}{z \left( 1 + \frac{1}{n+m} \text{tr} \Delta_j \mathbf{T}(z) \right)} \end{aligned} \quad (8)$$

and

$$\begin{aligned} \mathbf{T}(z) &= \left( \Psi(z)^{-1} - z \mathbf{\Gamma} \tilde{\Psi}(z) \mathbf{\Gamma}^H \right)^{-1} \\ \tilde{\mathbf{T}}(z) &= \left( \tilde{\Psi}(z)^{-1} - z \mathbf{\Gamma}^H \Psi(z) \mathbf{\Gamma} \right)^{-1} . \end{aligned} \quad (9)$$

Then this system admits a unique solution  $(\Psi_1(z), \dots, \Psi_N(z), \tilde{\Psi}_1(z), \dots, \tilde{\Psi}_{n+m}(z)) \in \mathcal{S}^{N+n+m}$ . In particular,  $\mathbf{T}(z)$  satisfies (7) for some measure  $\mu$ . Taking advantage of the particular forms of  $\rho_{ij}$  and  $\Delta_j$ , one can prove that  $\mathbf{T}(z)$  as defined in the previous equation satisfies

$$\mathbf{T}(z) = \left( -z \mathbf{I}_N + \mathbf{A} \mathbf{A}^H + \frac{1}{n} \sum_{j=1}^n \frac{\mathbf{D}_j}{1 + \frac{1}{n} \text{tr} \mathbf{D}_j \mathbf{T}(z)} \right)^{-1} .$$

Hence, the existence of a solution  $\mathbf{T}(z)$  to (6) is established; moreover  $\mathbf{T}(z)$  admits the representation (7).

To complete the proof of (i), it remains to check that such a  $\mathbf{T}(z)$  is unique. Assume that there exists  $\mathbf{T}(z)$  satisfying (6) with representation (7). Define  $\tilde{\Psi}(z)$  with the help of the

second part of (8),  $\Psi(z)$  with the help of the first part of (9) and  $\tilde{\mathbf{T}}(z)$  with the help of the second part of (9). It is then a matter of routine to check that  $\Psi(z)$  and  $\tilde{\Psi}(z)$  satisfy the system (8)-(9) (it remains basically to check that the first part of (8) is satisfied). As  $\mathbf{T}(z)$  admits the representation (7),  $\Psi_i(z)$  and  $\tilde{\Psi}_j(z)$  belong to  $\mathcal{S}$ . Hence  $\Psi(z)$  and  $\tilde{\Psi}(z)$  are uniquely defined and so is  $\mathbf{T}(z)$ .

Part (ii) of the theorem is a direct application of [6, Theorem 4.1]; details are therefore omitted.  $\blacksquare$

### III. FLUCTUATIONS OF $\mathcal{I}(\rho)$ : A CENTRAL LIMIT THEOREM

A number of studies has been devoted to the fluctuations of the mutual information, with various statistical assumptions for the channel  $\mathbf{H}$ ; see for instance [4], [13], and in a more mathematical flavor [14] (separable variance profile), [8] (general variance profile) and [15] (Rician channel with separable variance profile). A common feature of these works, although perhaps not much known, is the nice and concise closed-form expression of the variance of the mutual information which always writes

$$\Theta^2 = -\log \det (\mathbf{I} - \mathbf{J})$$

where  $\mathbf{J}$  is a Jacobian matrix associated to the set of fundamental equations of the matrix model under study. The fluctuations of the model  $\mathbf{H} \mathbf{H}^H + \mathbf{A} \mathbf{A}^H$  have not been studied yet, but relying on the previous observation, it is easy to infer the formula for the variance. Let  $\delta_j = \frac{1}{n} \text{tr} \mathbf{D}_j \mathbf{T}_\rho$ . Multiplying  $\mathbf{T}_\rho$  in (6) by  $\mathbf{D}_j$  and taking the normalized trace yields the following system of  $n$  equations:

$$\begin{aligned} \delta_j &= \frac{1}{n} \text{tr} \mathbf{D}_j \left( \rho \mathbf{I}_N + \mathbf{A} \mathbf{A}^H + \frac{1}{n} \sum_{k=1}^n \frac{\mathbf{D}_k}{1 + \delta_k} \right)^{-1} \\ &\triangleq \Gamma_j(\delta_1, \dots, \delta_n) . \end{aligned}$$

The computation of the  $n \times n$  Jacobian matrix  $\mathbf{J}_n$  of the function  $\Gamma = (\Gamma_1, \dots, \Gamma_n)$  is then straightforward:

$$[\mathbf{J}_n]_{k\ell} = \frac{\partial \Gamma_k}{\partial \delta_\ell} = \frac{1}{n} \frac{\frac{1}{n} \text{tr} \mathbf{D}_k \mathbf{T}_\rho \mathbf{D}_\ell \mathbf{T}_\rho}{(1 + \delta_\ell)^2} .$$

Based on the previous remarks, we are now in position to state the claim related to the fluctuations of the mutual information for the channel model under investigation.

*Claim 1 (The CLT):* Assume that Assumptions (A1), (A2) and (A3) hold true. Recall the definition of  $\mathbf{T}_\rho = \mathbf{T}(-\rho)$  and consider the following  $n \times n$  matrix  $\mathbf{J}_n$  defined by:

$$[\mathbf{J}_n]_{k,\ell} = \frac{1}{n} \frac{\frac{1}{n} \text{tr} \mathbf{D}_k \mathbf{T}_\rho \mathbf{D}_\ell \mathbf{T}_\rho}{\left( 1 + \frac{1}{n} \text{tr} \mathbf{D}_\ell \mathbf{T}_\rho \right)^2} . \quad (10)$$

Then:

(i) The real number  $\Theta_n^2 = -\log \det (\mathbf{I}_n - \mathbf{J}_n)$  is well-defined and satisfies

$$0 < \liminf_{N, n \rightarrow \infty} \Theta_n^2 \leq \limsup_{N, n \rightarrow \infty} \Theta_n^2 < \infty .$$

(ii) The following convergence holds true

$$\frac{N}{\Theta_n} (\mathcal{I}(\rho) - V(\rho)) \xrightarrow[N, n \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, 1)$$

where  $\mathcal{D}$  stands for the convergence in distribution.

The proof of part (i) closely follows [8, Theorem 3.1] and is therefore omitted. Due to the term  $\mathbf{A}\mathbf{A}^H$  in the model  $\mathbf{H}\mathbf{H}^H + \mathbf{A}\mathbf{A}^H$ , the proof of the fluctuations (ii) is not a simple consequence of Theorems 3.2 and 3.3 in [8] and necessitates special mathematical developments. We however provide the proof of the fluctuations in two specific cases, namely:

- 1) The case where  $\mathbf{A}\mathbf{A}^H = \mathbf{\Lambda}$  is a  $N \times N$  diagonal matrix.
- 2) The case where the variance profile is separable, i.e.,

$$\sigma_{ij} = \sqrt{d_i \tilde{d}_j}.$$

Beyond the proof of the fluctuations for these cases, simulations are provided that suggest the exactness of the variance formula in the general case.

*Proof:* [Proof of Claim 1 in case 1)] Let  $\mathbf{A}\mathbf{A}^H = \mathbf{\Lambda} = \text{diag}(\lambda_i^2; 1 \leq i \leq N)$ . Denote  $\mathbf{\Delta} = (\mathbf{\Lambda} + \rho \mathbf{I}_N)^{-1}$ , then:

$$\begin{aligned} \log \det(\mathbf{H}\mathbf{H}^H + \mathbf{\Lambda} + \rho \mathbf{I}_N) &= \\ - \log \det \mathbf{\Delta} + \log \det (\mathbf{\Delta}^{1/2} \mathbf{H}\mathbf{H}^H \mathbf{\Delta}^{1/2} + \mathbf{I}_N) &. \end{aligned}$$

Consider  $\tilde{\mathbf{H}} = \mathbf{\Delta}^{1/2} \mathbf{H}$ , then  $\tilde{\mathbf{H}}$  is a centered matrix with a variance profile given by:  $\kappa_{ij} = \sigma_{ij} / \sqrt{\lambda_i^2 + \rho}$ . Hence, the fluctuations of  $\log \det(\tilde{\mathbf{H}}\tilde{\mathbf{H}}^H + \mathbf{I}_N)$  fall into the framework of Theorems 3.2 and 3.3 in [8]. In particular,  $N(I(\rho) - V(\rho)) \rightarrow 0$  as  $N, n \rightarrow \infty$  and  $\tilde{\Theta}_n^{-1} N(\mathcal{I}(\rho) - I(\rho)) \rightarrow \mathcal{N}(0, 1)$  in distribution, where:

$$\tilde{\Theta}_n = -\log \det(\mathbf{I}_n - \tilde{\mathbf{J}}_n), \quad [\tilde{\mathbf{J}}_n]_{k,\ell} = \frac{1}{n} \frac{\frac{1}{n} \text{tr} \mathbf{D}_k \mathbf{\Upsilon} \mathbf{D}_\ell \mathbf{\Upsilon}}{(1 + \frac{1}{n} \text{tr} \mathbf{D}_\ell \mathbf{\Upsilon})^2}$$

and where  $\mathbf{\Upsilon}$  satisfies the following equation:

$$\mathbf{\Upsilon} = \left( \mathbf{I}_N + \frac{1}{n} \sum_{j=1}^n \frac{\mathbf{\Delta}_j}{1 + \frac{1}{n} \text{tr} \mathbf{\Delta}_j \mathbf{\Upsilon}} \right)^{-1} \quad \text{with} \quad \mathbf{\Delta}_j = \mathbf{D}_j \mathbf{\Delta}. \quad (11)$$

In order to establish Claim 1 in this case, it remains to prove that  $\tilde{\Theta}_n$  as just defined is equal to  $\Theta_n$ . From (11), it is straightforward to prove that  $\mathbf{\Delta} \mathbf{\Upsilon}$  satisfies (6) with  $z = -\rho$ , and is thus equal to  $\mathbf{T}$  due to the uniqueness of the solution of (6). It readily follows that  $\mathbf{J}_n = \tilde{\mathbf{J}}_n$ , which implies  $\Theta_n = \tilde{\Theta}_n$ . Claim 1 is proved in the case where  $\mathbf{A}\mathbf{A}^H = \mathbf{\Lambda}$ . ■

*Proof:* [Proof of Claim 1 in case 2)] In the case where the variance profile is separable, i.e.,  $\sigma_{ij} = \sqrt{d_i \tilde{d}_j}$ ,  $\mathbf{H}$  writes  $\mathbf{H} = n^{-1/2} \mathbf{D}^{1/2} \mathbf{W} \tilde{\mathbf{D}}^{1/2}$ , where  $\mathbf{D} = \text{diag}(d_i, 1 \leq i \leq N)$  and  $\tilde{\mathbf{D}} = \text{diag}(\tilde{d}_j, 1 \leq j \leq n)$  and where  $\mathbf{W}$  has i.i.d. standard complex Gaussian entries. Consider the following extended model:  $\tilde{\mathbf{W}} = [\mathbf{W} \mathbf{W}_1]$ , where  $\mathbf{W}_1$  is a  $N \times m$  matrix with i.i.d. standard complex Gaussian entries;  $\mathbf{\Delta} = \alpha \mathbf{D}$ , where  $\alpha = \sqrt{\frac{n+m}{n}}$ ;  $\mathbf{\Gamma} = [\mathbf{0}_{N \times n} \mathbf{A}]$ ; and finally  $\tilde{\mathbf{\Delta}} = \alpha \text{diag}(\tilde{\mathbf{D}}, \mathbf{0}_{m \times m})$ .

Then  $\mathbf{H}\mathbf{H}^H + \mathbf{A}\mathbf{A}^H$  writes

$$\left( \mathbf{\Delta}^{1/2} \frac{\tilde{\mathbf{W}}}{\sqrt{n+m}} \tilde{\mathbf{\Delta}}^{1/2} + \mathbf{\Gamma} \right) \left( \mathbf{\Delta}^{1/2} \frac{\tilde{\mathbf{W}}}{\sqrt{n+m}} \tilde{\mathbf{\Delta}}^{1/2} + \mathbf{\Gamma} \right)^H.$$

The CLT of the mutual information associated to this model has recently been established in [15]:

$$\frac{N}{\tilde{\Theta}_n} (\mathcal{I}(\rho) - I(\rho)) \xrightarrow[N, n \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, 1).$$

Moreover, it has been proved in [16, Theorem 2] that  $N(I(\rho) - V(\rho)) \rightarrow 0$ . Let us first provide the equations associated to this model in order to describe the variance  $\tilde{\Theta}_n^2$ . The following system in  $(\delta, \tilde{\delta})$  admits a unique pair of nonnegative solutions  $(\delta > 0, \tilde{\delta} > 0)$  (see for instance [16, Theorem 1]):

$$\begin{aligned} \delta &= \frac{1}{n+m} \text{tr} \mathbf{\Delta} \left( \rho(\mathbf{I}_N + \tilde{\delta} \mathbf{\Delta}) + \mathbf{\Gamma}(\mathbf{I}_{n+m} + \delta \tilde{\mathbf{\Delta}})^{-1} \mathbf{\Gamma}^H \right)^{-1} \\ \tilde{\delta} &= \frac{1}{n+m} \text{tr} \tilde{\mathbf{\Delta}} \left( \rho(\mathbf{I}_{n+m} + \delta \tilde{\mathbf{\Delta}}) + \mathbf{\Gamma}^H(\mathbf{I}_N + \tilde{\delta} \mathbf{\Delta})^{-1} \mathbf{\Gamma} \right)^{-1}. \end{aligned}$$

Introduce the matrices

$$\begin{aligned} \mathbf{\Upsilon} &= \left( \rho(\mathbf{I}_N + \tilde{\delta} \mathbf{\Delta}) + \mathbf{\Gamma}(\mathbf{I}_{n+m} + \delta \tilde{\mathbf{\Delta}})^{-1} \mathbf{\Gamma}^H \right)^{-1}, \\ \tilde{\mathbf{\Upsilon}} &= \left( \rho(\mathbf{I}_{n+m} + \delta \tilde{\mathbf{\Delta}}) + \mathbf{\Gamma}^H(\mathbf{I}_N + \tilde{\delta} \mathbf{\Delta})^{-1} \mathbf{\Gamma} \right)^{-1} \end{aligned}$$

and the quantities  $\gamma = (n+m)^{-1} \text{tr} \mathbf{\Delta}^2 \mathbf{\Upsilon}^2$  and  $\tilde{\gamma} = (n+m)^{-1} \text{tr} \tilde{\mathbf{\Delta}}^2 \tilde{\mathbf{\Upsilon}}^2$ . These quantities enable us to express the variance  $\tilde{\Theta}_n^2$  associated to the CLT as given in (12) at the top of the next page. Due to the particular form of the matrices associated to the extended model, one can readily prove that the variance takes the simpler form  $\tilde{\Theta}_n^2 = -\log(1 - \rho^2 \gamma \tilde{\gamma})$ .

It remains now to prove that  $\tilde{\Theta}_n = \Theta_n$ . Easy matrix computations yield to the fact that

$$\begin{aligned} \mathbf{\Upsilon} &= \left[ \rho \left( \mathbf{I} + \tilde{\delta} \mathbf{\Delta} \right) + \mathbf{A} \mathbf{A}^H \right]^{-1} \\ \tilde{\mathbf{\Upsilon}} &= \begin{bmatrix} \rho \left( \mathbf{I} + \alpha \delta \tilde{\mathbf{D}} \right) & \mathbf{0} \\ \mathbf{0} & \rho \mathbf{I} + \mathbf{A}^H \left( \mathbf{I} + \alpha \tilde{\delta} \mathbf{\Delta} \right) \mathbf{A} \end{bmatrix}^{-1} \end{aligned} \quad (13)$$

Hence, considering  $\tilde{\mathbf{\Upsilon}}$  as a block-diagonal matrix of inverses, we get:

$$\begin{aligned} \tilde{\delta} &= \frac{\alpha}{\rho(n+m)} \text{tr} \tilde{\mathbf{D}} \left[ \mathbf{I} + \alpha \delta \tilde{\mathbf{D}} \right]^{-1} \\ \tilde{\gamma} &= \frac{1}{n+m} \text{tr} \tilde{\mathbf{\Delta}}^2 \tilde{\mathbf{\Upsilon}}^2 = \frac{1}{\rho^2 n} \text{tr} \tilde{\mathbf{D}}^2 \left[ \mathbf{I} + \delta \tilde{\mathbf{D}} \alpha \right]^{-2}. \end{aligned} \quad (14)$$

Consider (6), which defines  $\mathbf{T}$ , for  $z = -\rho$ ; note that  $\mathbf{D}_j = \tilde{d}_j \mathbf{D}$  and introduce  $\kappa = \frac{1}{n} \text{tr} \mathbf{D} \mathbf{T}_\rho$  so that  $\mathbf{T}_\rho$  satisfies the equation:

$$\begin{aligned} \mathbf{T}_\rho &= \left[ \rho \mathbf{I} + \mathbf{A} \mathbf{A}^H + \frac{1}{n} \sum_{j=1}^n \frac{\tilde{d}_j \mathbf{D}}{1 + \kappa \tilde{d}_j} \right]^{-1} \\ &= \left[ \rho \mathbf{I} + \mathbf{A} \mathbf{A}^H + \left( \frac{1}{n} \text{tr} \tilde{\mathbf{D}} (\mathbf{I} + \kappa \tilde{\mathbf{D}})^{-1} \right) \mathbf{D} \right]^{-1}. \end{aligned}$$

$$\tilde{\Theta}_n^2 = -\log \left( \left( 1 - \frac{1}{m+n} \text{tr} \mathbf{\Delta}^{1/2} \mathbf{\Upsilon} \mathbf{\Gamma} (\mathbf{I} + \delta \tilde{\mathbf{\Delta}})^{-1} \tilde{\mathbf{\Delta}} (\mathbf{I} + \delta \tilde{\mathbf{\Delta}})^{-1} \mathbf{\Gamma}^H \mathbf{\Upsilon} \mathbf{\Delta}^{1/2} \right)^2 - \rho^2 \gamma \tilde{\gamma} \right) \quad (12)$$

Considering the definitions of  $\mathbf{\Upsilon}$  and  $\tilde{\delta}$  as given in (13) and (14), one can prove that  $\mathbf{T}_\rho$  and  $\mathbf{\Upsilon}$  satisfy the same equation and hence are equal. In particular,  $\kappa = \alpha\delta$  and  $\tilde{\Theta}_n^2$  writes:

$$\tilde{\Theta}_n^2 = -\log \left( 1 - \frac{1}{n} \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2 \times \frac{1}{n} \text{tr} \mathbf{D}^2 [\mathbf{I} + \kappa \tilde{\mathbf{D}}]^{-2} \right).$$

We now rewrite  $\Theta_n^2$  as given in Claim 1:

$$\mathbf{J}_{k\ell} = \frac{\frac{1}{n} \tilde{d}_k \tilde{d}_\ell \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2}{n (1 + \tilde{d}_\ell)^2}.$$

Hence,

$$\mathbf{J} = \left( \frac{1}{n^2} \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2 \right) \mathbf{u}^H \mathbf{v}$$

where  $\mathbf{u} = [\tilde{d}_1, \dots, \tilde{d}_n]$  and  $\mathbf{v} = [\tilde{d}_1(1 + \kappa \tilde{d}_1)^{-2}, \dots, \tilde{d}_n(1 + \kappa \tilde{d}_n)^{-2}]$ . Now,

$$\begin{aligned} & -\log \det \left( \mathbf{I} - \left( \frac{1}{n^2} \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2 \right) \mathbf{u}^H \mathbf{v} \right) \\ &= -\log \left( 1 - \left( \frac{1}{n^2} \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2 \right) \mathbf{v} \mathbf{u}^H \right) \\ &= -\log \left( 1 - \frac{1}{n} \text{tr} \mathbf{D}^2 \mathbf{T}_\rho^2 \times \frac{1}{n} \text{tr} \mathbf{D}^2 [\mathbf{I} + \kappa \tilde{\mathbf{D}}]^{-2} \right) \end{aligned}$$

which is exactly the expression of  $\tilde{\Theta}_n^2$  and the proof is completed. ■

#### IV. NUMERICAL RESULTS

In order to verify the accuracy of the analysis in the previous sections, we provide now some simulation results. We consider a variance profile where each  $\sigma_{ij}^2$  is drawn randomly from the interval  $[0, 10]$ . The interference covariance matrix  $\mathbf{A} \mathbf{A}^H$  is also generated in a random fashion by letting  $\mathbf{A} = \frac{1}{\sqrt{N}} \mathbf{X}$ , where  $\mathbf{X}$  is a standard complex Gaussian  $N \times m$  matrix. Both the variance profile ( $\sigma_{ij}^2$ ) and  $\mathbf{A}$  are chosen at random at the beginning of the simulations and then kept constant. We define the signal-to-noise-ratio as  $\text{SNR} = 1/\rho$  and let  $m = 3$ .

Fig. 1 shows the normalized ergodic mutual information  $I(\rho)$  versus the SNR for several different values of  $N$  and  $n$ . Solid lines represent the deterministic equivalent approximation  $V(\rho)$  as given by Theorem 1. Markers are obtained by Monte Carlo simulations for 10,000 different realizations of  $\mathbf{H}$ . We observe a very accurate fit between both results which demonstrates that the asymptotic analysis yields accurate approximations for small channel dimensions.

Fig. 2 and Fig. 3 depict the histogram of the random variable  $\frac{N}{\Theta_n} (\mathcal{I}(\rho) - V(\rho))$  in comparison with the normal distribution  $\mathcal{N}(0, 1)$  for two different pairs of parameters  $N, n$ . The overlap is almost perfect for a rather large system with  $N = 16$  receive antennas and  $n = 8$  transmitters (Fig. 3). In the case  $N = 2$  and  $n = 1$  (Fig. 2), the fluctuations

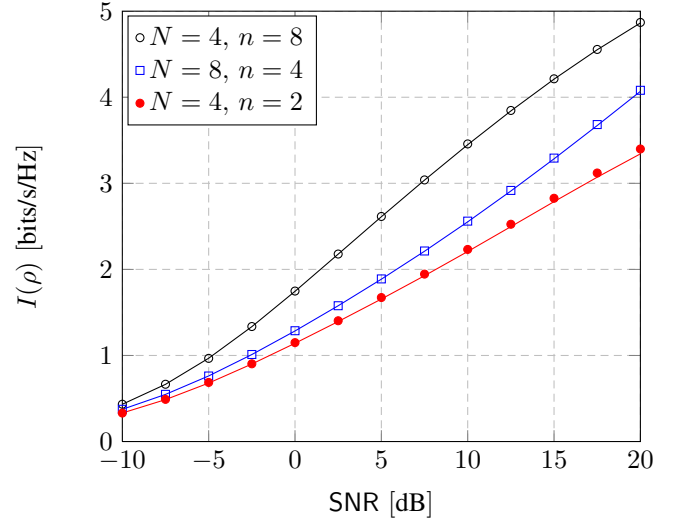


Fig. 1. Normalized ergodic mutual information  $I(\rho)$  versus SNR for different channel dimensions  $N, n$ . Solid lines correspond to the deterministic equivalent approximation  $V(\rho)$ . Markers are obtained by Monte Carlo simulations.

match also surprisingly well although there is clearly a bias. This bias has a simple explanation: While the fluctuations of  $N(\mathcal{I} - V)$  arise from those of  $N(\mathcal{I} - \mathbb{E}\mathcal{I}) = N(\mathcal{I} - I)$ , the deterministic difference  $N(I - V)$  yields a bias of order  $\mathcal{O}(N^{-1})$  (see for instance [16, Theorem 2]). Although this bias vanishes quickly with  $N$ , it is of order one, as noticed in Figure 2, for very small values of  $n, N$ . Hence, these plots further validate the CLT as stated in Claim 1.

#### V. CONCLUSIONS

We have studied the fluctuations of the mutual information of a class of large-dimensional MIMO channels with arbitrary colored noise. First, we have provided a deterministic approximation of the mutual information, which is tight in the asymptotic limit. Second, we have established the fluctuations of the mutual information around this approximation in form of a CLT. Both analytical results have then been confirmed by simulations and it was shown that the asymptotic results yield accurate approximations for even small channel dimensions. The results can be readily applied in various settings of interfering multi-user MIMO networks or distributed antenna systems.

#### VI. ACKNOWLEDGMENTS

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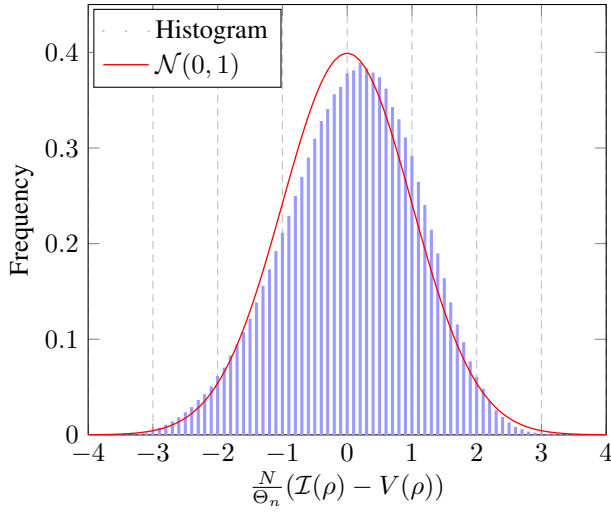


Fig. 2. Histogram of  $\frac{N}{\Theta_n}(\mathcal{I}(\rho) - V(\rho))$  in comparison with the normal distribution  $\mathcal{N}(0, 1)$  for  $N = 2$ ,  $n = 1$ .

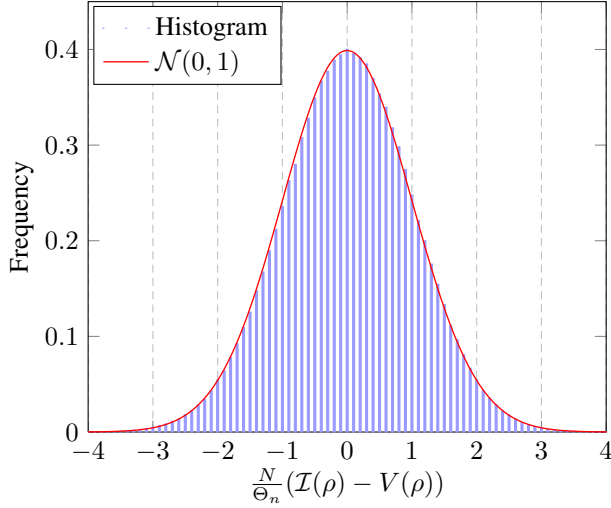


Fig. 3. Histogram of  $\frac{N}{\Theta_n}(\mathcal{I}(\rho) - V(\rho))$  in comparison with the normal distribution  $\mathcal{N}(0, 1)$  for  $N = 16$ ,  $n = 8$ .

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